



Model Order Reduction for Multi-scale,
Multi-physics Problems
Background

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Outline

Introduction (today)

Theory (today)

Practice (tomorrow)

The leading edge (tomorrow)



Theory
Outline

SVD, QR decompositions

Least squares regression

Sampling & Reconstruction

Sensing

Resources

<https://caslab.engin.umich.edu/teaching>

- Isaac Newton Institute tutorial on Model Order reduction for complex systems (Jan 2023)
 1. [Model Order Reduction theory manual](http://websites.umich.edu/~caslab/docs/Newton/MOR_Theory.pdf)
http://websites.umich.edu/~caslab/docs/Newton/MOR_Theory.pdf
 2. [PERFORM](#) (Prototyping environment for reacting flow order reduction methods : code)
 3. [PERFORM](#) (Prototyping environment for reacting flow order reduction methods : doc)
 4. Slides (coming soon)

Singular Value Decomposition

Any matrix has a singular value decomposition.

SVD of a general matrix $A \in \mathbb{C}^{m \times n}$ is given by

$$\underbrace{A}_{m \times n} = \underbrace{U}_{m \times m} \underbrace{\Sigma}_{m \times n} \underbrace{V^*}_{n \times n}$$

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- Singular values are typically arranged in descending order.

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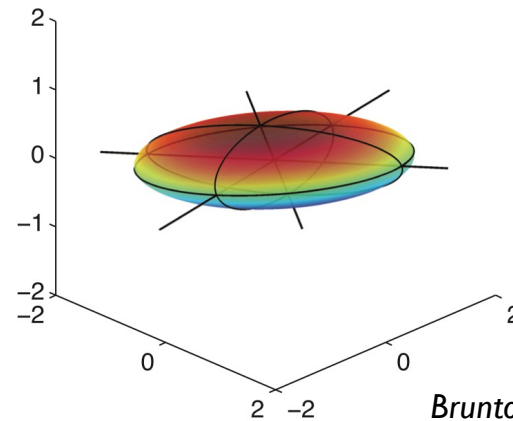
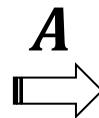
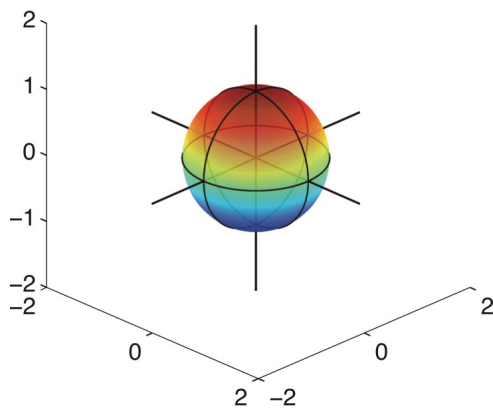
- Columns of U are the left singular vectors of A .
- Columns of V are the right singular vectors of A .
- The left and right singular vectors are orthonormal.
- U and V are unitary matrices.

Geometric Interpretation of SVD

Matrix multiplication introduces a rotation and a stretching action.

Singular values are the lengths of semi-axes of the hyper-ellipsoid obtained as a result of operation of matrix A on the unit hypersphere.

In 3-D:



Brunton and Kutz 2019

Singular Value Decomposition

Columns of A (snapshots matrix) can be

- Measurements from experiments
- Image pixels
- State of a physical system (velocity, pressure, etc.) at discrete points
- ...

Rank of A is equal to the number of non-zero singular values.

$$\underbrace{\mathbf{A}}_{m \times n} = \underbrace{\mathbf{U}}_{m \times m} \underbrace{\mathbf{\Sigma}}_{m \times n} \underbrace{\mathbf{V}^*}_{n \times n} \quad \rightarrow \quad \underbrace{\mathbf{A}}_{m \times n} = \underbrace{\hat{\mathbf{U}}}_{m \times r} \underbrace{\hat{\mathbf{\Sigma}}}_{r \times r} \underbrace{\hat{\mathbf{V}}^*}_{r \times n}$$

SVD Facts

$$A = U\Sigma V^*$$

- I. Every matrix has a SVD and singular values are uniquely determined, but they are not necessarily distinct.

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3. Singular values of A are the square roots of the eigenvalues of A^*A

$$\begin{aligned} A^*A &= [U\Sigma V^*]^* [U\Sigma V^*] \\ &= V\Sigma^2 V^*, \end{aligned}$$

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4. Same can be proved for the eigenvalues of AA^* .
5. The rank of A is equal to the number of its non-zero singular values.

6.
$$\|A\|_2 = \sqrt{\lambda_{\max}(A^*A)} = \sigma_{\max}(A)$$

SVD Facts

7. Given $\mathbf{A} = \sum_{j=1}^r \sigma_j \mathbf{u}_j \mathbf{v}_j^*$ for any $0 < q < r$ the matrix $\mathbf{A}_q = \sum_{j=1}^q \sigma_j \mathbf{u}_j \mathbf{v}_j^*$ satisfies the following properties:

$$\|\mathbf{A} - \mathbf{A}_q\|_2 = \sigma_{q+1}$$

$$\|\mathbf{A} - \mathbf{A}_q\|_F = \sqrt{\sum_{i=q+1}^r \sigma_i^2}$$

$$\|\mathbf{A}^+ - \mathbf{A}_q^+\|_F = \sqrt{\sum_{i=q+1}^r \frac{1}{\sigma_i^2}}$$

Image Compression with SVD

$$\underbrace{\mathbf{A}}_{m \times n} = \underbrace{\mathbf{U}}_{m \times m} \underbrace{\mathbf{\Sigma}}_{m \times n} \underbrace{\mathbf{V}^*}_{n \times n}$$



$$\underbrace{\mathbf{A}}_{m \times n} = \underbrace{\hat{\mathbf{U}}}_{m \times r} \underbrace{\hat{\mathbf{\Sigma}}}_{r \times r} \underbrace{\hat{\mathbf{V}}^*}_{r \times n}$$



Full Image

10 Modes



20 Modes

30 Modes



40 Modes

100 Modes





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QR Factorization

If $A \in \mathbb{C}^{m \times n}$ and A has full column rank, then the QR decomposition is given by

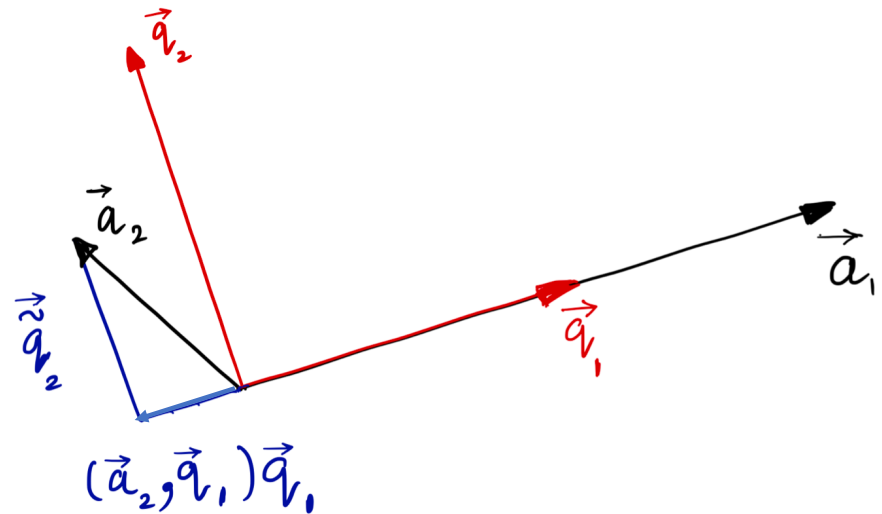
$$A = QR$$

- $Q \in \mathbb{C}^{m \times n} = \{\mathbf{q}_1, \mathbf{q}_2, \dots, \mathbf{q}_n\}$ are orthonormal vectors.
- $R \in \mathbb{C}^{n \times n}$ is an upper-triangular matrix with non-zero diagonal elements.
- QR factorization can be computed by the Gram-Schmidt procedure.

Gram-Schmidt Procedure

Given a linearly independent set of vectors $S \equiv \{\mathbf{a}_1, \mathbf{a}_2, \mathbf{a}_3, \dots, \mathbf{a}_n\}$, where $\mathbf{a}_i \in \mathbb{R}^m$, the Gram-Schmidt procedure finds an orthonormal set of vectors $\{\mathbf{q}_1, \mathbf{q}_2, \mathbf{q}_3, \dots, \mathbf{q}_n\}$ that spans the same subspace as S . The procedure is as follows:

1. $\tilde{\mathbf{q}}_1 = \mathbf{a}_1$
2. $\mathbf{q}_1 = \frac{\tilde{\mathbf{q}}_1}{\|\tilde{\mathbf{q}}_1\|}$
3. $\tilde{\mathbf{q}}_2 = \mathbf{a}_2 - (\mathbf{q}_1^T \mathbf{a}_2)\mathbf{q}_1$
4. $\mathbf{q}_2 = \frac{\tilde{\mathbf{q}}_2}{\|\tilde{\mathbf{q}}_2\|}$
5. $\tilde{\mathbf{q}}_3 = \mathbf{a}_3 - (\mathbf{q}_1^T \mathbf{a}_3)\mathbf{q}_1 - (\mathbf{q}_2^T \mathbf{a}_3)\mathbf{q}_2$
6. $\mathbf{q}_3 = \frac{\tilde{\mathbf{q}}_3}{\|\tilde{\mathbf{q}}_3\|}$
7. Repeat recursively



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To compute QR factorization of A using the Gram-Schmidt procedure:

$$\mathbf{Q} = [\mathbf{q}_1 \quad \mathbf{q}_2 \quad \dots \quad \mathbf{q}_n]$$

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To compute QR factorization of A using the Gram-Schmidt procedure:

$$\mathbf{Q} = [\mathbf{q}_1 \quad \mathbf{q}_2 \quad \dots \quad \mathbf{q}_n]$$

$$\mathbf{R} = \begin{bmatrix} \|\tilde{\mathbf{q}}_1\| & \mathbf{q}_1^T \mathbf{a}_2 & \dots & \mathbf{q}_1^T \mathbf{a}_n \\ 0 & \|\tilde{\mathbf{q}}_2\| & \dots & \mathbf{q}_2^T \mathbf{a}_n \\ 0 & 0 & \dots & \mathbf{q}_i^T \mathbf{a}_n \\ \vdots & \vdots & \dots & \vdots \\ 0 & 0 & \dots & \|\tilde{\mathbf{q}}_n\| \end{bmatrix}$$



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Least-squares Regression

Consider an over-determined system of equations

$$\mathbf{Ax} = \mathbf{y}$$

$$\mathbf{A} = \begin{bmatrix} A_{11} & A_{12} & \cdots & A_{1n} \\ A_{21} & A_{22} & \cdots & A_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ A_{m1} & A_{m2} & \cdots & A_{mn} \end{bmatrix}, \quad \mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{bmatrix}, \quad \mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_m \end{bmatrix}$$

Least-squares Regression

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$$m > n$$

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y: Observables (data, snapshots of quantities of interest)

x: Model parameters (**unknown**)

Least-squares Regression

$$\begin{matrix} & A & & x & = & b \\ \left[\begin{array}{c} \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \end{array} \right] & & \left[\begin{array}{c} \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \end{array} \right] & = & \left[\begin{array}{c} \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \\ \text{gray rectangle} \end{array} \right] \end{matrix}$$

Least-squares Regression

$$\mathbf{Ax} = \mathbf{y}$$

An over-determined system usually has no solution.

💡 But we can search for parameters that fit the equations best.

This can be done by solving the following optimization problem:

$$\hat{\mathbf{x}} = \mathit{arg} \min_{\mathbf{x}} S(\mathbf{x})$$

Objective function:

$$S(\mathbf{x}) = \|\mathbf{y} - \mathbf{Ax}\|_2^2$$

Least-squares Regression

$$\mathbf{Ax} = \mathbf{y}$$

The optimization problem has a unique solution if A is full column rank.

$$\hat{\mathbf{x}} = \mathit{arg} \min_{\mathbf{x}} S(\mathbf{x})$$



$$S(\mathbf{x}) = \|\mathbf{y} - \mathbf{Ax}\|_2^2 = (\mathbf{y} - \mathbf{Ax})^T (\mathbf{y} - \mathbf{Ax}) = \mathbf{y}^T \mathbf{y} - \mathbf{x}^T \mathbf{A}^T \mathbf{y} - \mathbf{y}^T \mathbf{Ax} + \mathbf{x}^T \mathbf{A}^T \mathbf{Ax}$$

Substitute $(\mathbf{x}^T \mathbf{A}^T \mathbf{y})^T = \mathbf{y}^T \mathbf{Ax}$

$$S(\mathbf{x}) = \mathbf{y}^T \mathbf{y} - 2\mathbf{x}^T \mathbf{A}^T \mathbf{y} + \mathbf{x}^T \mathbf{A}^T \mathbf{Ax}$$

Least-squares Regression

$$\mathbf{Ax} = \mathbf{y}$$

$$S(\mathbf{x}) = \mathbf{y}^T \mathbf{y} - 2\mathbf{x}^T \mathbf{A}^T \mathbf{y} + \mathbf{x}^T \mathbf{A}^T \mathbf{A} \mathbf{x}$$

Differentiate with respect to \mathbf{x} :

$$-\mathbf{A}^T \mathbf{y} + (\mathbf{A}^T \mathbf{A}) \mathbf{x} = 0 \quad \Rightarrow \quad \text{First-order condition}$$

$\mathbf{A}^T \mathbf{A}$: The Gramian matrix of \mathbf{A}

$\mathbf{A}^T \mathbf{y}$: The moment matrix

Least-squares Regression

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First-order condition

The solution to the optimization problem:

$$\hat{\mathbf{x}} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{y} = \mathbf{A}^+ \mathbf{y}$$

The **second-order condition** for the minimum:

$$\mathbf{A}^T \mathbf{A} > \mathbf{0}$$

Pseudoinverse (from last lecture)

Any matrix has a pseudoinverse.

But A^+ can only be computed explicitly under these conditions:

I. If A is full column rank, then $A^T A$ is invertible and $A^+ = (A^* A)^{-1} A^*$.

- This is called the **left inverse** as in this case $A^+ A = I$.
- In this case the pseudoinverse can also be computed by the QR decomposition:

$$A^+ = R^{-1} Q^T$$

Least-squares Regression

$$\mathbf{Ax} = \mathbf{y}$$

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$$\hat{\mathbf{x}} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{y} = \mathbf{A}^+ \mathbf{y}$$

$\mathbf{A}\hat{\mathbf{x}}$ is the orthogonal projection of \mathbf{y} onto the *range*(\mathbf{A}):

$$\mathcal{P}_{\text{range}(\mathbf{A})}(\mathbf{y}) = \mathbf{A}\hat{\mathbf{x}} = \mathbf{A}(\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{y}$$

Least-squares Regression

$$\mathbf{Ax} = \mathbf{y}$$

These statements are equivalent:

- $\mathbf{A}\hat{\mathbf{x}}$ is the orthogonal projection of \mathbf{y} onto the $\text{range}(\mathbf{A})$:

$$\mathcal{P}_{\text{range}(\mathbf{A})}(\mathbf{y}) = \mathbf{A}\hat{\mathbf{x}} = \mathbf{A}(\mathbf{A}^T\mathbf{A})^{-1}\mathbf{A}^T\mathbf{y}$$

- The residual $\mathbf{r} = \mathbf{A}\hat{\mathbf{x}} - \mathbf{y}$ is orthogonal to the range of \mathbf{A} .

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- \mathbf{r} is in the null space of \mathbf{A}^T
- $\mathbf{A}^T(\mathbf{A}\hat{\mathbf{x}} - \mathbf{y}) = 0$
- $(\mathbf{r}, \mathbf{Az}) = 0 \quad \forall \mathbf{z} \in \mathbb{R}^n$

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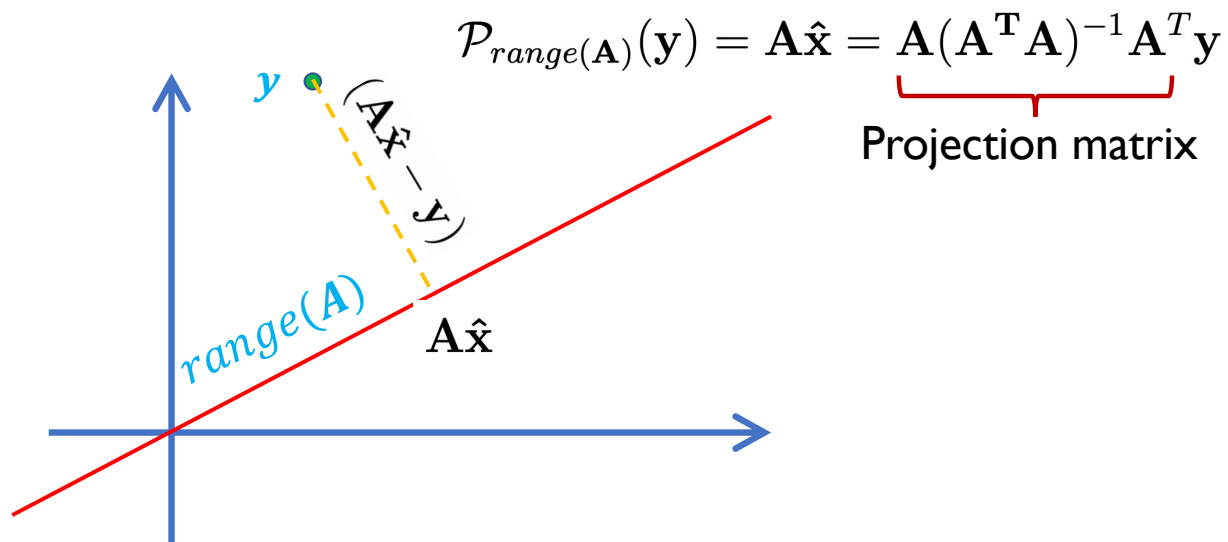
$$\mathcal{P}_{\text{range}(\mathbf{A})}(\mathbf{y}) = \mathbf{A}\hat{\mathbf{x}} = \underbrace{\mathbf{A}(\mathbf{A}^T\mathbf{A})^{-1}\mathbf{A}^T}_{\text{Projection matrix}}\mathbf{y}$$

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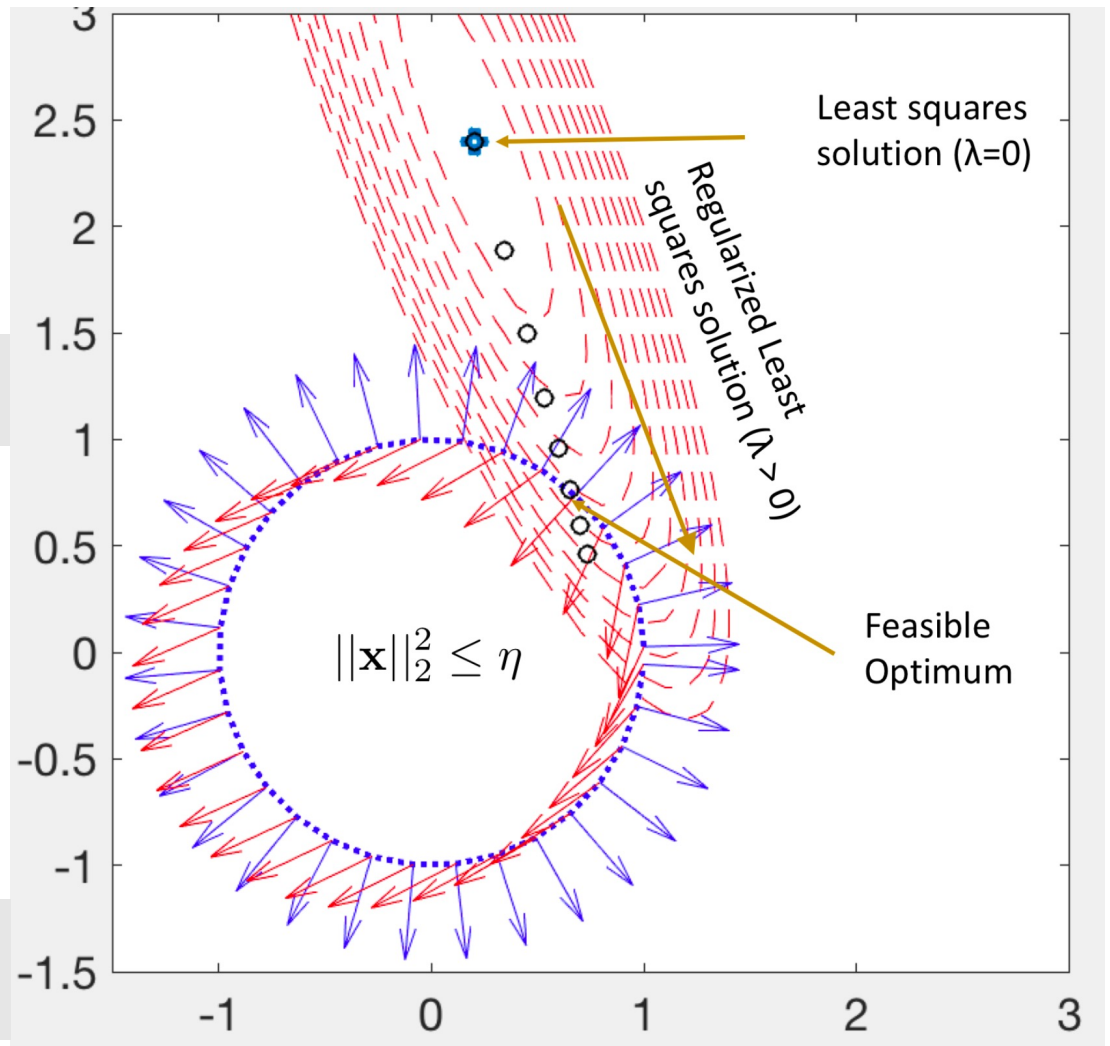
If $\mathbf{y} \in \text{range}(\mathbf{A})$ then we can satisfy $\mathbf{y} = \mathbf{A}\hat{\mathbf{x}}$ precisely.

Otherwise, we can satisfy $\mathcal{P}_{\text{range}(\mathbf{A})}(\mathbf{y}) = \mathbf{A}\hat{\mathbf{x}}$.

What about regularized Least Squares?

$$\min_{\mathbf{x}_1} \mathcal{F}(\mathbf{x}_1, \lambda) = \|\mathbf{A}\mathbf{x}_1 - \mathbf{y}\|_2^2 + \lambda\|\mathbf{x}_1\|_2^2$$

$$\min_{\mathbf{x}_2} \|\mathbf{A}\mathbf{x}_2 - \mathbf{y}\|_2^2 \quad s.t. \quad \|\mathbf{x}_2\|_2^2 \leq \eta$$



Consider $p(y|\theta) = \mathcal{N}(y; A\theta, C)$ with $A \in \mathbb{R}^{n \times m}$. Assume $p(\theta) = \mathcal{N}(\theta; \theta_0, B)$. Thus, $Y \sim \mathcal{N}(A\theta_0, ABA^T + C)$
 To mimic a linear regression problem, consider the goal of estimating

$$\theta^* = \arg \min_{\theta} L(\theta)$$

What about SGD?

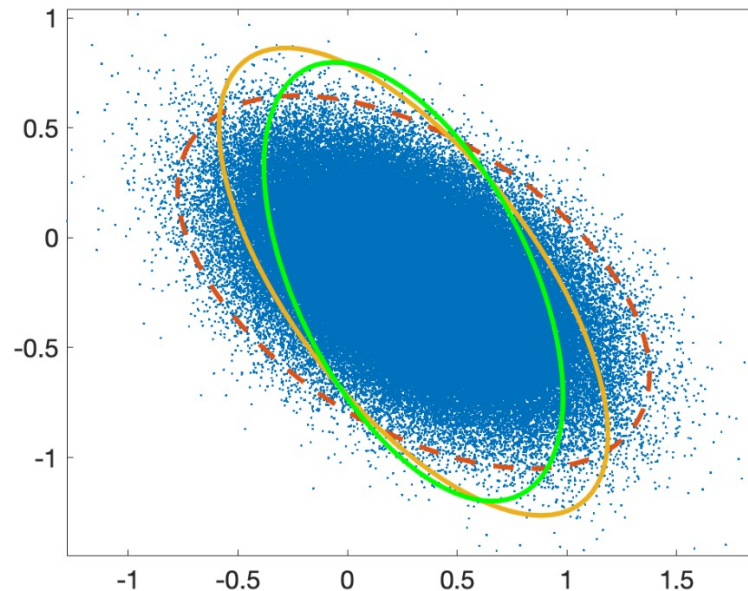
where

$$L(\theta) = \mathbb{E}_Y[(y - A\theta)^T C^{-1}(y - A\theta)] + (\theta - \theta_0)^T B^{-1}(\theta - \theta_0)$$

SGD Converges to:

$$\theta_{\mu} = [A^T C^{-1} A + B^{-1}]^{-1} (A^T C^{-1} \mu_Y + B^{-1} \theta_0)$$

$$\begin{aligned} \theta_{\Sigma} &= \alpha [A^T C^{-1} A + B^{-1}]^{-1} A^T C^{-1} J J^T C^{-1} A \\ &= \alpha [A^T C^{-1} A + B^{-1}]^{-1} A^T C^{-1} \Sigma_Y C^{-1} A \end{aligned}$$





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SVD, QR decompositions

Least squares regression

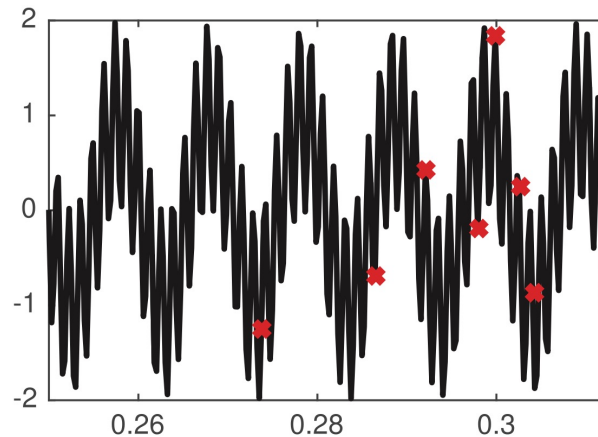
Sampling & Reconstruction

Sensing

Compression, Sensing and Reconstruction

It is not always possible/efficient to collect high-dimensional measurement data.

Is it possible to measure the quantity of interest at a few sensor locations and use these measurements to reconstruct the entire signal?



Sampling and Reconstruction

Let's consider a signal $\mathbf{x} \in \mathbb{R}^m$, represented by basis $\Psi \in \mathbb{R}^{m \times n}$ and basis coefficients $\mathbf{a} \in \mathbb{R}^n$,

$$\mathbf{x} = \Psi \mathbf{a}$$

Sampling and Reconstruction

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Goal:

Instead of using the entire signal \mathbf{x} , we want to subsample \mathbf{x} , and see if we can reconstruct \mathbf{x} using sparse measurements.

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$$\mathbf{x} = \Psi \mathbf{a}$$

In order to subsample \mathbf{x} , let's define a matrix $\mathbf{P} \in \mathbb{R}^{p \times m}$, so that,

$$\mathbf{y} = \mathbf{P} \mathbf{x} \in \mathbb{R}^p$$

is the signal obtained by subsampling \mathbf{x} .

Sampling and Reconstruction

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Example:

If $\mathbf{x} \in \mathbb{R}^5$, and we want to sample only the fourth and second measurements of \mathbf{x} , then matrix \mathbf{P} takes the form,

$$\mathbf{P} = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 \\ 0 & 1 & 0 & 0 & 0 \end{bmatrix}$$

Sampling and Reconstruction

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Therefore,

$$\mathbf{P} \mathbf{x} = \mathbf{y} = \mathbf{P} \Psi \mathbf{a}$$

Sampling and Reconstruction

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$$\mathbf{x} = \Psi \mathbf{a}$$

We subsample \mathbf{x} by,

$$\mathbf{P}\mathbf{x} = \mathbf{y} = \mathbf{P}\Psi\mathbf{a}$$

Therefore, the basis coefficients \mathbf{a} can be estimated as,

$$\hat{\mathbf{a}} = [\mathbf{P}\Psi]^+ \mathbf{y}$$

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And reconstruct \mathbf{x} as,

$$\hat{\mathbf{x}} = \Psi[\mathbf{P}\Psi]^+ \mathbf{y}$$

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Note:

If the original signal had p non-zero coefficients, then p measurements reconstruct the signal exactly.

Sampling and Reconstruction

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And reconstruct \mathbf{x} as,

$$\hat{\mathbf{x}} = \Psi [\mathbf{P} \Psi]^+ \mathbf{y}$$

Questions:

- How do we know $\hat{\mathbf{a}}$ is sparse?
- What if we have noisy measurements?
- How to choose optimal sensor locations?

Sampling and Reconstruction

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- What if we have noisy measurements?
- How to choose optimal sensor locations?
- How many measurements do we collect?

Sampling and Reconstruction

Questions:

- How do we know $\hat{\mathbf{a}}$ is sparse?

Sparse basis coefficients can be found by solving,

$$\hat{\mathbf{a}} = \min_{\mathbf{a}} \|\mathbf{a}\|_1 \text{ such that } \mathbf{y} = \mathbf{P}\Psi\mathbf{a}$$

This optimization problem is convex.

Sampling and Reconstruction

Questions:

- What if we have noisy measurements?

$$\mathbf{y} = \mathbf{P}\Psi\mathbf{a} + \epsilon$$

$$\epsilon \sim \mathcal{N}(0, \sigma^2)$$

Sparse basis coefficients can be found by solving,

$$\hat{\mathbf{a}} = \min_{\mathbf{a}} \|\mathbf{a}\|_1 \text{ such that } \|\mathbf{y} - \mathbf{P}\Psi\mathbf{a}\|_2 \leq \sigma$$

Sampling and Reconstruction

Questions:

- How to choose optimal sensor locations? (how to define matrix P ?)

Rule:

Rows of P have to be orthogonal to the columns of Ψ .

$$\mathbf{P}\mathbf{x} = \mathbf{y} = \mathbf{P}\Psi\mathbf{a}$$

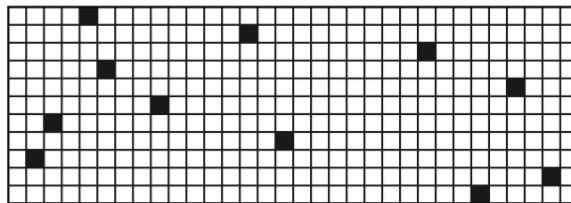
Sampling and Reconstruction

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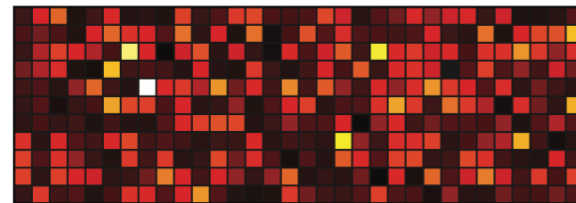
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Possible choices for matrix P ,

(a) Random single pixel



(b) Gaussian random



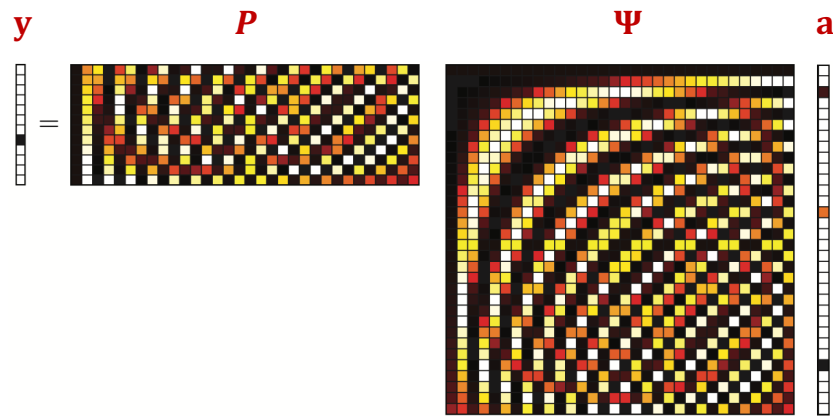
Brunton and Kutz, 2019

Sampling and Reconstruction

Questions:

- How to choose optimal sensor locations? (how to define matrix P ?)

An example of a bad choice for matrix P ,



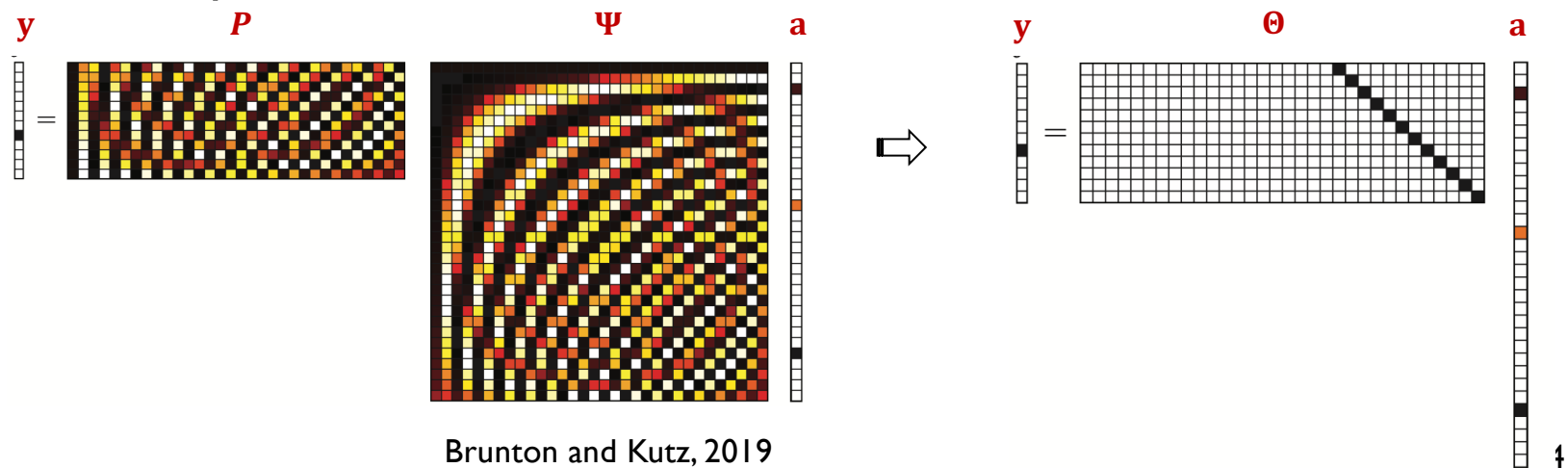
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Sampling and Reconstruction

Questions:

- How to choose optimal sensor locations? (how to define matrix P ?)

An example of a bad choice for matrix P ,



Sampling and Reconstruction

Questions:

- How many measurements do we collect?

Theorem:

According to the Shannon-Nyquist sampling theorem, in order to recover a signal we should sample the signal twice the rate of its highest frequency.



Outline

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Sensing

Sensing: Empirical Interpolation

Goal:

We want to design a sparse measurement matrix \mathbf{P} such that the inverse problem is as well-conditioned as possible.

$$\mathbf{y} = \mathbf{P}\Psi\mathbf{a} \quad \mathbf{a} \in \mathbb{R}^n$$

Sensing: Empirical Interpolation

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In general, we may use randomly placed sensors to estimate \mathbf{a} . However, **when $p = n$** (the number of measurements equals the number of modes), $\mathbf{P}\Psi$ is often numerically singular.

Sensing: Empirical Interpolation

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We want to design a sparse measurement matrix \mathbf{P} such that the inverse problem is as well-conditioned as possible. $\mathbf{y} = \mathbf{P}\Psi\mathbf{a}$ $\mathbf{a} \in \mathbb{R}^n$

In general, we may use randomly placed sensors to estimate \mathbf{a} . However, **when $p = n$** (the number of measurements equals the number of modes), $\mathbf{P}\Psi$ is often numerically singular.

Solution:

- Oversampling
- Using QR factorization

Sensing: Empirical Interpolation

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We want to design a sparse measurement matrix \mathbf{P} such that the inverse problem is as well-conditioned as possible.

$$\mathbf{y} = \mathbf{P}\Psi\mathbf{a}$$

Idea:

If we know the type of the signal, it is possible to design optimized sensors using low-rank features extracted from patterns in the data.

Given a basis Ψ , we can find a \mathbf{P}^* by solving,

$$\mathbf{P}^* = \min_{\mathbf{P}} \|\mathbf{x} - \Psi[\mathbf{P}\Psi]^+ \mathbf{y}\|_2$$

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$$\mathbf{P}^* = \min_{\mathbf{P}} \|\mathbf{x} - \underbrace{\Psi[\mathbf{P}\Psi]^+}_{\text{Reconstruction of } \mathbf{x} \text{ using sparse basis coefficients}} \mathbf{y}\|_2$$

Reconstruction of \mathbf{x} using
sparse basis coefficients

Sampling and Reconstruction

Let's consider a signal $\mathbf{x} \in \mathbb{R}^m$, represented by basis $\Psi \in \mathbb{R}^{m \times n}$ and basis coefficients $\mathbf{a} \in \mathbb{R}^n$,

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If $p = n$, then,

$$\mathbf{P}^* = \min_{\mathbf{P}} \|\mathbf{x} - \Psi[\mathbf{P}\Psi]^{-1} \mathbf{y}\|_2$$

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Given a basis Ψ , we can find a \mathbf{P}^* by solving,

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For an orthonormal basis Ψ ,

$$\|\mathbf{x} - \Psi[\mathbf{P}\Psi]^{-1}\mathbf{P}\mathbf{x}\|_2 \leq \|[\mathbf{P}\Psi]^{-1}\|_2 \|\mathbf{I} - \Psi\Psi^T\|_2 \|\mathbf{x}\|_2$$

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Projection error

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For an orthonormal basis Ψ ,

$$\|\mathbf{x} - \Psi[\mathbf{P}\Psi]^{-1}\mathbf{P}\mathbf{x}\|_2 \leq \underbrace{\|[\mathbf{P}\Psi]^{-1}\|_2}_{\text{Sampling error}} \underbrace{\|[\mathbf{I} - \Psi\Psi^T]\mathbf{x}\|_2}_{\text{Projection error}}$$

Projection error

Sensing: Discrete Empirical Interpolation

Given a basis Ψ , we can find a \mathbf{P}^* by solving,

$$\mathbf{P}^* = \min_{\mathbf{P}} \|\mathbf{x} - \Psi[\mathbf{P}\Psi]^{-1}\mathbf{y}\|_2$$

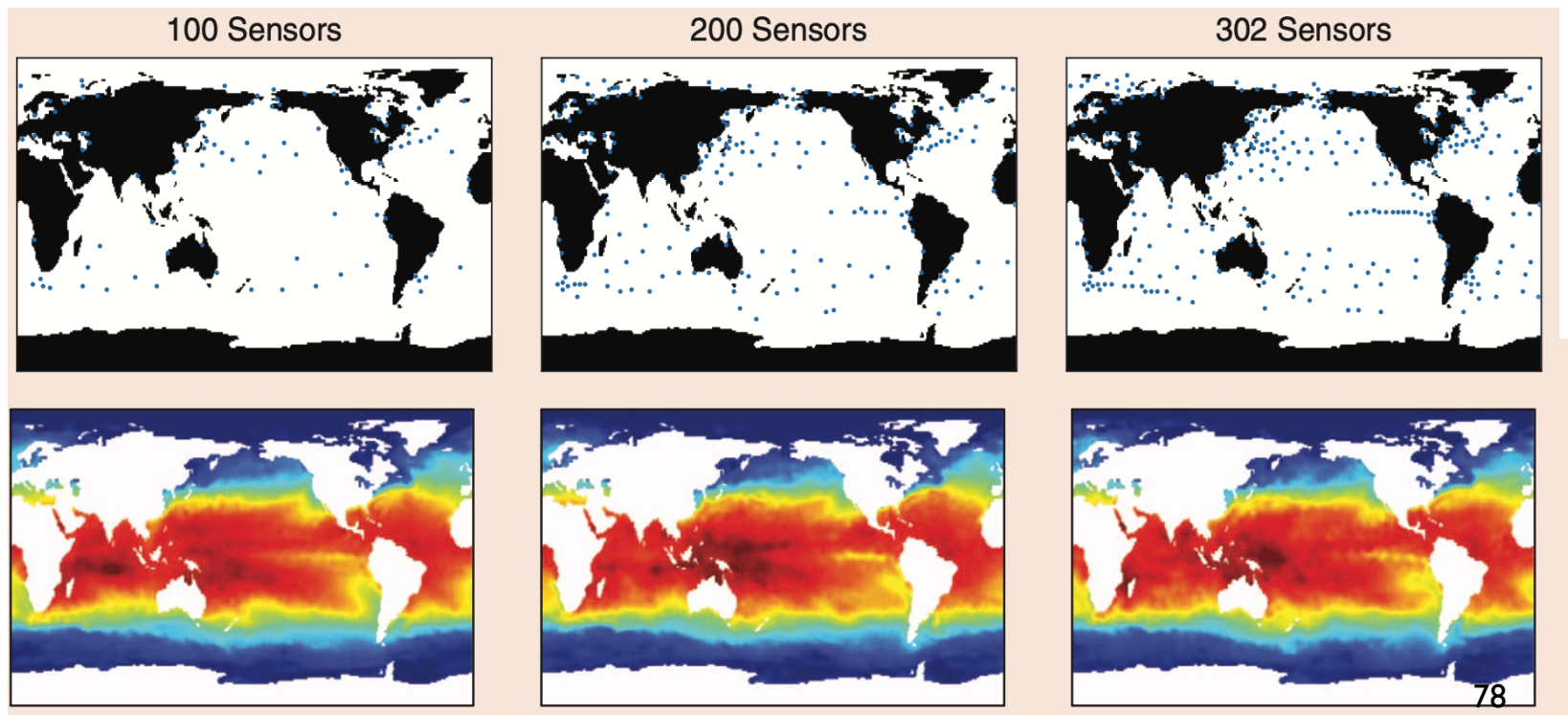
For an orthonormal basis Ψ ,

$$\|\mathbf{x} - \Psi[\mathbf{P}\Psi]^{-1}\mathbf{P}\mathbf{x}\|_2 \leq \|[\mathbf{P}\Psi]^{-1}\|_2 \|[\mathbf{I} - \Psi\Psi^T]\mathbf{x}\|_2$$

Therefore, the sensor selection process reduces to,

$$\mathbf{P}^* = \min_{\mathbf{P}} \|[\mathbf{P}\Psi]^{-1}\|_2$$

Sea Surface temperature dataset (Manohar et al.)



Sensing: QDEIM

The QDEIM approach uses QR factorization to obtain the sampling matrix (sensor locations).

Why QR factorization?

Reminder: we want to design a sparse measurement matrix \mathbf{P} such that the inverse problem is as well-conditioned as possible.

$$\mathbf{y} = \mathbf{P}\Psi\mathbf{a}$$

QR factorization with pivoting, maximizes the submatrix volume and controls the condition number.

Sensing: QDEIM

Side Note: QR factorization with column pivoting

QR decomposition with column pivoting decomposes a matrix into a unitary matrix \mathbf{Q} , an upper triangular matrix \mathbf{R} , and a column permutation matrix Φ ,

$$\mathbf{W}\Phi = \mathbf{Q}\mathbf{R}$$

Where, $\mathbf{Q} \in \mathbb{C}^{n \times n}$ and $\mathbf{R} \in \mathbb{C}^{n \times m}$.

Sensing: QDEIM

Side Note: QR factorization with column pivoting

Given $\mathbf{W} \in \mathbb{R}^{n \times m}$, with a column pivoted QR decomposition we have,

$$\mathbf{W}\Phi = \mathbf{Q}\mathbf{R}$$

Where, $\mathbf{Q} \in \mathbb{C}^{n \times n}$ and $\mathbf{R} \in \mathbb{C}^{n \times m}$, therefore,

$$\mathbf{W}\Phi = \mathbf{Q}[\mathbf{R}_1 \ \mathbf{R}_2]$$

$\mathbf{R}_1 \in \mathbb{C}^{n \times n}$ is an upper triangular matrix.

The columns of \mathbf{W} are permuted such that the diagonal elements of \mathbf{R}_1 are non-increasing.

Sensing: QDEIM

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$$\mathbf{W}\Phi = \mathbf{Q}[\mathbf{R}_1 \ \mathbf{R}_2]$$

$\mathbf{R}_1 \in \mathbb{C}^{n \times n}$ is an upper triangular matrix.

Also,

$$\sigma_i^2 = |r_{ii}|^2 \ ; \ 1 \leq i \leq n$$

Sensing: QDEIM

The QDEIM approach uses QR factorization to obtain the sampling matrix (sensor locations).

$$\mathbf{W}\Phi = \mathbf{Q}\mathbf{R}$$

Let's set $\mathbf{W} = \Psi^T$, and

$$\mathbf{P}\Psi = [\mathbf{Q}\mathbf{R}_1]^T = \mathbf{R}_1^T \mathbf{Q}^T$$

We want to solve

$$\mathbf{P}^* = \min_{\mathbf{P}} \|\mathbf{P}\Psi\|_2^{-1}$$

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$$\mathbf{P}\Psi = [\mathbf{Q}\mathbf{R}_1]^T = \mathbf{R}_1^T \mathbf{Q}^T$$

We want to solve

$$\mathbf{P}^* = \min_{\mathbf{P}} \|\mathbf{P}\Psi\|_2^{-1}$$

Therefore,

$$\|\mathbf{P}\Psi\|_2 = \|\mathbf{R}_1^T \mathbf{Q}^T\|_2 = \sigma_{max}(\mathbf{R}_1)$$

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Therefore,

$$\|\mathbf{P}\Psi\|_2 = \|\mathbf{R}_1^T \mathbf{Q}^T\|_2 = \sigma_{max}(\mathbf{R}_1)$$

and,

$$\|\mathbf{P}\Psi\|_2^{-1} = \frac{1}{\sigma_{min}(\mathbf{R}_1)}$$

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and,

$$\|\mathbf{P}\Psi\|_2^{-1} = \frac{1}{\sigma_{min}(\mathbf{R}_1)}$$

We should keep $\sigma_{min}(\mathbf{R}_1)$ as large as possible.

Sensing: QDEIM

The QDEIM approach uses QR factorization to obtain the sampling matrix (sensor locations).

We want to minimize sampling error,

$$\|[\mathbf{P}\Psi]^{-1}\|_2 = \frac{1}{\sigma_{\min}(\mathbf{R}_1)}$$

We should keep $\sigma_{\min}(\mathbf{R}_1)$ as large as possible.

QR factorization with column pivoting expands the submatrix volume by enforcing a diagonal dominance structure (Manohar, 2018),

$$\sigma_i^2 = |r_{ii}|^2 \geq \sum_{j=i}^k |r_{jk}|^2; \quad 1 \leq i \leq k \leq m$$

Sensing: QDEIM

The QDEIM approach uses QR factorization to obtain the sampling matrix (sensor locations).

Let's set $\mathbf{W} = \Psi^T$ in

$$\mathbf{W}\Phi = \mathbf{Q}[\mathbf{R}_1 \ \mathbf{R}_2]$$

QDEIM solves the column pivoted QR factorization

$$\Psi^T \Phi = \mathbf{Q}\mathbf{R}_1$$

and the optimal sampling matrix is,

$$\mathbf{P}^* = \Phi^T$$

Sensing: QDEIM

The QDEIM approach uses QR factorization to obtain the sampling matrix \mathbf{C} .

```
if (p==r) % QR sensor selection, p=r
    [Q,R,pivot] = qr(Psi_r', 'vector');  $\Psi_r^T \mathbf{C}^T = \mathbf{QR}$ 
elseif (p>r) % Oversampled QR sensors, p>r
    [Q,R,pivot] = qr(Psi_r*Psi_r', 'vector');  $(\Psi_r \Psi_r^T) \mathbf{C}^T = \mathbf{QR}$ 
end
C = zeros(p,n);
for j=1:p
    C(j,pivot(j))=1;
end
```

Brunton and Kutz, 2019

p : number of sensors

r : number of basis functions $\mathbf{a} \in \mathbb{R}^r$